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ON THE PROBABILITY FUNCTIONAL OF DIFFUSION PROCESSES*

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It is known that the multivariate probability density of a Markov diffusion process $x(t)$, $0 \leq t \leq T$, described by the equation

$$(1) \quad \frac{\partial p(x, t)}{\partial t} = - \frac{\partial}{\partial x} [a(x, t)p] + \frac{1}{2} \frac{\partial^2}{\partial x^2} [b(x, t)p]$$

can be written approximately in the form

$$(2) \quad p(x_0, x_1, \dots, x_T) = p_0(x_0) \prod_{i=0}^{N-1} [2\pi b(x_i, t_i)]^{-\frac{1}{2}} \\ \times \exp \left\{ -\frac{1}{2} \sum_{i=0}^{N-1} \left[\frac{x_{i+1} - x_i}{\Delta_i} - a(x_i, t_i) \right]^2 \frac{\Delta_i}{b(x_i, t_i)} \right\},$$

where $x_i = x(t_i)$; $t_{i+1} - t_i = \Delta_i > 0$; $t_N = T$, $t_0 = 0$.

The smaller $\Delta = \max [\Delta_0, \dots, \Delta_{N-1}]$, the higher the accuracy of the above formula. For small Δ , the summation in the exponent recalls, by its form, the Darboux sum corresponding to the integral

$$(3) \quad -\frac{1}{2} \int_0^T [\dot{x} - a(x, t)]^2 \frac{dt}{b(x, t)}, \quad (x = x(t))$$

(Here and in the sequel a dot denotes the time derivative.)

It is natural to inquire whether one can assign some precise significance to such an integral, and not just a symbolic one.

The realizations of $x(t)$ almost certainly do not have finite derivatives and, a fortiori, the latter are not square-integrable. Moreover the consideration of functionals of the type (3) is rather interesting from the point of view of applications, since in practice, as a rule, the realizations of a diffusion process are not exactly Markov, but smooth ones with a finite derivative. For such processes an

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integral of the type (3) has an exact meaning.

One can conduct a systematic study of such smooth processes and their functionals by an explicit introduction to the theory of the operation of smoothing. However a simpler approach is also of interest, one without an explicit consideration of smoothing but which deals with functionals of smooth functions. For such functions one can take, within a known approximation, observed smoothed realizations.

A useful step in that direction is the introduction of the probability functional $W[x(t)]$ defined on the space B of functions $x(t)$ having a bounded continuous derivative $\dot{x}(t)$ of bounded variation.

1. Let $z(t)$ be a Wiener process with initial condition $z(0) = z_0$, described by the equation

$$(4) \quad \frac{\partial p(z, t)}{\partial t} = \frac{1}{2} \frac{\partial^2 p(z, t)}{\partial z^2}.$$

The multivariate density (2) is defined in this case by the exact equality

$$(5) \quad p(z_1, \dots, z_N | z_0) = \text{const} \exp \left\{ -\frac{1}{2} \sum_{i=1}^{N-1} \left(\frac{z_{i+1} - z_i}{\Delta_i} \right)^2 \Delta_i \right\}.$$

It is natural to define the probability functional by the formula

$$(6) \quad W[z(t)] = \exp \left\{ -\frac{1}{2} \int_0^T [\dot{z}(t)]^2 dt \right\}.$$

Both in this case and in more general cases, the probability functional is defined only up to an arbitrary (finite) constant factor. We shall choose this factor in such a way as to obtain the simplest possible expression.

Let the functions $z(t) \in B$ for which the functional (6) is defined fulfill the conditions

$$(7) \quad |\dot{z}(t)| < M_z, \quad 0 \leq t \leq T;$$

$$(8) \quad \sum_k |\dot{z}(\tau_{k+1}) - \dot{z}(\tau_k)| < M_z.$$

(Here $\dots < \tau_k < \tau_{k+1} < \dots$ are points at which $z(t)$ takes extremal values.) Condition (8) may be replaced by the inequality

(9)

Such replacement is always derivative \dot{z} and its integral are

THEOREM. 1. Let $z^{(1)}(t)$ and $z^{(2)}(0) = z_0$, $h(t) > 0$, $0 \leq t \leq$

$$(10) \quad \lim_{\epsilon \rightarrow 0} \frac{P\{z^{(1)} < z^{(2)}\}}{P\{z^{(1)} < z^{(2)}\}}$$

PROOF. Let us make the

$$(11) \quad \bar{z}(t) =$$

The last process is described by the

$$\frac{\partial p(\bar{z}, t)}{\partial t} =$$

According to the results of continuous, and the corresponding

$$\frac{d\mu_{\bar{z}}[z(t)]}{d\mu_z[z(t)]} = \exp \left\{ \int [\dot{\bar{z}}(t) - \dot{z}(t)]^2 dt \right\}$$

$$(12) \quad = \exp \left\{ -\frac{1}{2} \int_0^T [\dot{\bar{z}}(t) - \dot{z}(t)]^2 dt \right\}$$

where

$$I = \int_0^T [\dot{\bar{z}}(t) - \dot{z}(t)]^2 dt$$

the Radon-Nikodým theorem with

$$(13) \quad \mu_{\bar{z}}(\lambda)$$

where

$$\Lambda_\epsilon = \left\{ z(t) : z^{(2)} < z(t) \right\}$$

Substituting (12) into (13), we

$$(9) \quad \int_0^T |z| dt < M'_z.$$

Such replacement is always permissible without further restrictions if the derivative \dot{z} and its integral are understood in the generalized sense.

THEOREM. 1. Let $z^{(1)}(t), z^{(2)}(t), h(t)$ belong to B , and let $z^{(1)}(0) = z^{(2)}(0) = z_0, h(t) > 0, 0 \leq t \leq T$ and $\epsilon > 0$. Then

$$(10) \quad \lim_{\epsilon \rightarrow 0} \frac{P \{ z^{(1)} < z < z^{(1)} + \epsilon h, 0 \leq t \leq T \}}{P \{ z^{(2)} < z < z^{(2)} + \epsilon h, 0 \leq t \leq T \}} = \frac{W[z^{(1)}]}{W[z^{(2)}]}.$$

PROOF. Let us make the change of variables

$$(11) \quad \tilde{z}(t) = z(t) + z^{(2)}(t) - z^{(1)}(t).$$

The last process is described by the equation

$$\frac{\partial p(\tilde{z}, t)}{\partial t} = -(\dot{z}^{(2)} - \dot{z}^{(1)}) \frac{\partial p}{\partial \tilde{z}} + \frac{1}{2} \frac{\partial^2 p}{\partial \tilde{z}^2}.$$

According to the results of [1] the processes $z(t)$ and $\tilde{z}(t)$ are absolutely continuous, and the corresponding functional derivative is equal to

$$(12) \quad \begin{aligned} \frac{d\mu_{\tilde{z}}[z(t)]}{d\mu_z[z(t)]} &= \exp \left\{ \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}] dz(t) - \frac{1}{2} \int_0^T [z^{(2)} - z^{(1)}]^2 dt \right\} \\ &= \exp \left\{ -\frac{1}{2} \int_0^T \dot{z}^{(1)2} dt + \frac{1}{2} \int_0^T \dot{z}^{(2)2} dt + I \right\}, \end{aligned}$$

where

$$I = \int_0^T [z^{(2)} - z^{(1)}] d[z - z^{(2)}].$$

By the Radon-Nikodým theorem we have

$$(13) \quad \mu_z(\Lambda_\epsilon) = \int_{\Lambda_\epsilon} \frac{d\mu_{\tilde{z}}}{d\mu_z} d\mu_z.$$

Here

$$\Lambda_\epsilon = \left\{ z(t) : z^{(2)} < z(t) < z^{(2)} + \epsilon h, 0 \leq t \leq T; z(0) = z_0 \right\}.$$

Substituting (12) into (13), we find

$$(14) \quad \frac{P \{ z^{(1)} < z < z^{(1)} + \epsilon h, 0 \leq t \leq T \}}{P \{ z^{(2)} < z < z^{(2)} + \epsilon h, 0 \leq t \leq T \}} = \frac{\mu_z(\Lambda_\epsilon)}{\mu_z(\Lambda_\epsilon)}$$

$$= \exp \left\{ -\frac{1}{2} \int_0^T z^{(1)2} dt + \frac{1}{2} \int_0^T z^{(2)2} dt \right\} \frac{1}{\mu_z(\Lambda_\epsilon)} \int_{\Lambda_\epsilon} e^t d\mu_z.$$

In the expression

$$I = \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}] d[z - z^{(2)}] = [\dot{z}_T^{(2)} - \dot{z}_T^{(1)}] [z_T - z_T^{(2)}] - \int_0^T [z - z^{(2)}] [\ddot{z}^{(2)} - \ddot{z}^{(1)}] dt$$

one can, according to inequalities (7), (9), carry out the following estimates

$$|z - z^{(2)}| < \epsilon h < \epsilon M_h, \quad 0 \leq t \leq T \quad \text{for } z(t) \in \Lambda_\epsilon;$$

$$\int |z - z^{(2)}| |\ddot{z}^{(2)} + \ddot{z}^{(1)}| dt < \epsilon M_h \int |\ddot{z}^{(2)} - \ddot{z}^{(1)}| dt$$

$$\leq \epsilon M_h \int [|\dot{z}^{(2)}| + |\dot{z}^{(1)}|] dt < \epsilon M_h (M'_{z^{(2)}} + M'_{z^{(1)}});$$

$$|\dot{z}_T^{(2)} - \dot{z}_T^{(1)}| < M'_{z^{(2)}} + M'_{z^{(1)}},$$

from which follows

$$|I| < \epsilon M_h [M'_{z^{(1)}} + M'_{z^{(2)}} + M'_{z^{(1)}} + M'_{z^{(2)}}] \equiv \epsilon M,$$

for all $z(t) \in \Lambda_\epsilon$. Therefore $|e' - 1| < \epsilon M + \frac{1}{2} \epsilon^2 M^2 + \dots$ and

$$\lim_{\epsilon \rightarrow 0} \frac{1}{\mu_z(\Lambda_\epsilon)} \int_{\Lambda_\epsilon} (e' - 1) d\mu_z = 0,$$

i. e.

$$(15) \quad \lim_{\epsilon \rightarrow 0} \frac{1}{\mu_z(\Lambda_\epsilon)} \int_{\Lambda_\epsilon} e' d\mu_z = 1.$$

As a consequence of this and of (14), after passing to the limit as $\epsilon \rightarrow 0$, (10) follows.

2. Consider the Markov diffusion process $z(t)$, $0 < t < T$, corresponding to the equation

$$(16) \quad \frac{\partial p(z, t)}{\partial t} = -\frac{\partial}{\partial z} [m(z, t) p] + \frac{1}{2} \frac{\partial^2 p}{\partial z^2},$$

where $m(z, t)$ is a bounded function with respect to z . The probability satisfies the same equation (16), but $p(z, t) = p(z - z_1, t)$. Let us transform this equation defined by the equality

$$(17) \quad p_{t,t}(z_1, z) =$$

where $f(z, t)$ is a function which will be determined. Calculations lead to the equation

$$(18) \quad \frac{\partial \bar{p}}{\partial t} = \frac{1}{2} \frac{\partial^2 \bar{p}}{\partial z^2} - \frac{\partial}{\partial z} \left[\left(m + \frac{\partial f}{\partial z} \right) \bar{p} \right]$$

Choose the function $f(z, t)$ such that \bar{p} is zero, i. e. set

$$(19) \quad \frac{\partial f(z, t)}{\partial z} = m(z, t),$$

The solution of the resulting equation

$$(20) \quad \frac{\partial \bar{p}}{\partial t} = \frac{1}{2} \frac{\partial^2 \bar{p}}{\partial z^2} +$$

may be written, as we know, in the form

$$(21) \quad \bar{p}_{t,t}(z_1, z_2) = \int_{C_{z_1, z_2}} \exp \left\{ -\frac{1}{2} \int_0^t \dots \right\}$$

with respect to the conditional Wiener measure. Here and in the sequel

$$C_{z_1, \dots, z_k} = \left\{ z(t) : \dots \right\}$$

According to (17), the multivariate probability

$$p(z_1, \dots, z_N | z_0) = P_{0,t}(z_1, \dots, z_N | z_0)$$

can be written in the form

$$p(z_1, \dots, z_T | z_0) = e^{f(z_1, \dots, z_T) - f(z_0)}$$

if we take (21) into account,

where $m(z, t)$ is a bounded function with uniformly continuous first derivative with respect to z . The probability density $p_{t_1 t}(z_1, z)$ of jump from z_1 to z satisfies the same equation (16), but with initial condition $p_{t_1 t_1}(z_1, z) = \delta(z - z_1)$. Let us transform this equation, going over to the function $\tilde{p}_{t_1 t}(z_1, z)$ defined by the equality

$$(17) \quad p_{t_1 t}(z_1, z) = e^{f(z, t)} \tilde{p}_{t_1 t}(z_1, z) e^{-f(z, t)},$$

where $f(z, t)$ is a function which will be defined explicitly in the sequel. Direct calculations lead to the equation

$$(18) \quad \frac{\partial \tilde{p}}{\partial t} = \frac{1}{2} \frac{\partial^2 \tilde{p}}{\partial z^2} - \frac{\partial}{\partial z} \left[\left(m + \frac{\partial f}{\partial z} \right) \tilde{p} \right] + \left[m \frac{\partial f}{\partial z} + \frac{\partial f}{\partial t} + \frac{1}{2} \frac{\partial^2 f}{\partial z^2} + \frac{1}{2} \left(\frac{\partial f}{\partial z} \right)^2 \right] \tilde{p}.$$

Choose the function $f(z, t)$ such that the term with first derivative reduces to zero, i. e. set

$$(19) \quad \frac{\partial f(z, t)}{\partial z} = m(z, t), \quad f(z, t) = - \int_0^z m(z', t) dz'.$$

The solution of the resulting equation

$$(20) \quad \frac{\partial \tilde{p}}{\partial t} = \frac{1}{2} \frac{\partial^2 \tilde{p}}{\partial z^2} + \left[\frac{\partial f}{\partial t} - \frac{1}{2} m^2 - \frac{1}{2} \frac{\partial m}{\partial z} \right] \tilde{p}$$

may be written, as we know, in the form of the functional integral

$$(21) \quad \tilde{p}_{t_1 t_2}(z_1, z_2) = \int_{C_{z_1, z_2}} \exp \left\{ - \frac{1}{2} \int_{t_1}^{t_2} \left[m^2 + \frac{\partial m}{\partial z} - 2 \frac{\partial f}{\partial t} \right] dt \right\} dw [z | z_0]$$

with respect to the conditional Wiener measure $w[\Lambda | z_1] = P(\Lambda | z(t_1) = z_1)$.

Here and in the sequel

$$C_{z_1, \dots, z_k} = \left\{ z(t) : z(t_i) = z_i, \quad i = 1, \dots, k \right\}.$$

According to (17), the multivariate distribution

$$p(z_1, \dots, z_N | z_0) = p_{0 t_1}(z_0, z_1) \dots p_{t_{N-1} t_N}(z_{N-1}, z_N)$$

can be written in the form

$$p(z_1, \dots, z_T | z_0) = e^{f(z_0, 0) - f(z_T, T)} \tilde{p}_{0 t_1}(z_0, z_1) \dots \tilde{p}_{t_{N-1} T}(z_{N-1}, z_T).$$

if we take (21) into account,

$$\begin{aligned}
 & p(z_1, \dots, z_T | z_0) = e^{f(z_0, 0) - f(z_T, T)} \times \\
 (22) \quad & \int_{c_{z_1, \dots, z_T}} \exp \left\{ -\frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} - 2 \frac{\partial f}{\partial t} \right] dt \right\} dw[z | z_0].
 \end{aligned}$$

Integrating the last expression, we find

$$\begin{aligned}
 \mu[\Lambda | z_0] &= \int_{\Lambda} \exp \left\{ f(z_0, 0) - f(z_T, T) \right. \\
 (23) \quad & \left. - \frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} - 2 \frac{\partial f}{\partial t} \right] dt \right\} dw[z | z_0],
 \end{aligned}$$

where

$$\Lambda = \Lambda_{t_1, \dots, t_N} \left\{ z(t) : z(t_i) \in E_i, \quad i=1, \dots, N \right\}.$$

Since in this formula $N, t_1, \dots, t_N, E_1, \dots, E_N$ are arbitrary, equation (23) holds for an arbitrary set Λ of more general form by virtue of the separability and continuity of the process under consideration. Therefore the measures $\mu[\Lambda | z_0]$ and $w[\Lambda | z_0]$ are absolutely continuous, and the corresponding functional derivative is equal to

$$(24) \quad \frac{d\mu[z | z_0]}{dw[z | z_0]} = \exp \left\{ f(z_0, 0) - f(z_T, T) + \int_0^T \frac{\partial f}{\partial t} dt - \frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} \right] dt \right\}.$$

Taking into account that $d\mu[z] = P(dz_0) d\mu[z | z_0]$ and assuming the existence of an initial probability density $p_0(z_0) = P(dz_0) / dz_0$, we have

$$\begin{aligned}
 d\mu[z] &= p_0(z_0) \exp \left\{ f(z_0, 0) - f(z_T, T) + \int_0^T \frac{\partial f}{\partial t} dt \right. \\
 (25) \quad & \left. - \frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} \right] dt \right\} dw[z | z_0] dz_0.
 \end{aligned}$$

Since the probability functional (6) corresponds to the Wiener measure $w[\Lambda | z_0]$, it is natural then, as is seen from (25), to define the probability functional of the given process $z(t)$ by the formula

$$\begin{aligned}
 W[z(t)] &= p_0(z_0) \exp \\
 & - \frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} - 2 \frac{\partial f}{\partial t} \right] dt
 \end{aligned}$$

Theorem 1 in turn holds for the process in the same fashion similar to that of the previous section. After substitution of the probability functional (25) into (21) we obtain

$$(26) \quad \frac{\partial p(\bar{z}, t)}{\partial t} = -\frac{\partial}{\partial \bar{z}} \left\{ \left[m(\bar{z}, t) \frac{\partial p(\bar{z}, t)}{\partial \bar{z}} + \Phi(\bar{z}, t) p(\bar{z}, t) \right] \right\}$$

Applying formula (24) to this expression we obtain

$$\begin{aligned}
 \frac{d\mu_z[\bar{z} | \bar{z}_0]}{dw[\bar{z} | \bar{z}_0]} &= \exp \left\{ f(z_0, 0) - f(z_T, T) \right. \\
 & + \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}] d\bar{z} - \frac{1}{2} \int_0^T m^2(\bar{z}, t) dt \\
 (27) \quad & \left. - \int_0^T m(\bar{z}, t) [\dot{z}^{(2)} - \dot{z}^{(1)}] dt \right\}
 \end{aligned}$$

Expressing $d\mu_z[\bar{z}] / d\mu_z(\bar{z})$ as the ratio of the derivatives

$$\begin{aligned}
 \frac{d\mu_z[\bar{z}]}{d\mu_z(\bar{z})} &= \frac{p_0(\bar{z}_0 - z_0^{(2)} + z_0^{(1)})}{p_0(\bar{z}_0)} \exp \\
 & + \int_0^T \Phi(\bar{z} - z^{(2)} + z^{(1)}) dt \\
 (28) \quad & - \frac{1}{2} \int_0^T \dot{z}^{(1)2} dt + \frac{1}{2} \int_0^T \dot{z}^{(2)2} dt
 \end{aligned}$$

where

$$\Phi(z) = -\frac{1}{2} m^2(z)$$

If in (29) one replaces the functional derivative $d\mu_z[\bar{z}] / d\mu_z(\bar{z})$ by $(\bar{z}_0 - z_0^{(2)} + z_0^{(1)}) \ln p_0(\bar{z}_0), \dots, p_0(z_0^{(2)}), \dots$, expression (29) becomes

$$(26) \quad W[z(t)] = p_0(z_0) \exp \left\{ f(z_0, 0) - f(z_T, T) + \int_0^T \frac{\partial f}{\partial t} dt - \frac{1}{2} \int_0^T \left[m^2 + \frac{\partial m}{\partial z} \right] dt - \frac{1}{2} \int_0^T z^2 dt \right\}.$$

Theorem 1 in turn holds for this functional. Its proof can be carried out in a fashion similar to that of the previous proof, with some unessential complications. After substitution of the process (11), the equation for $z(t)$ takes the form

$$(27) \quad \frac{\partial p(\bar{z}, t)}{\partial t} = - \frac{\partial}{\partial \bar{z}} \left\{ m(\bar{z} - z^{(2)} + z^{(1)}, t) + \dot{z}^{(2)} - \dot{z}^{(1)} \right\} p + \frac{1}{2} \frac{\partial^2 p}{\partial \bar{z}^2}.$$

Applying formula (24) to this case, we find

$$(28) \quad \frac{d\mu_{\bar{z}}[\bar{z} | \bar{z}_0]}{d\mu[\bar{z} | \bar{z}_0]} = \exp \left\{ f(z_0, 0) - f(z_T, T) + \int_0^T \left[\frac{\partial f(z, t)}{\partial t} + \frac{\partial f(z, t)}{\partial z} (-\dot{z}^{(2)} + \dot{z}^{(1)}) \right] dt + \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}] d\bar{z} - \frac{1}{2} \int_0^T \left[m^2(z, t) + \frac{\partial m(z, t)}{\partial z} \right] dt - \int_0^T m(z, t) [\dot{z}^{(2)} - \dot{z}^{(1)}] dt - \frac{1}{2} \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}]^2 dt \right\}$$

with

$$z = \bar{z} - z^{(2)} + z^{(1)}.$$

Expressing $d\mu_{\bar{z}}[\bar{z}] / d\mu_z(\bar{z})$ as the ratio of (28) to (24), we find

$$(29) \quad \frac{d\mu_{\bar{z}}[\bar{z}]}{d\mu_z[\bar{z}]} = \frac{p_0(\bar{z}_0 - z_0^{(2)} + z_0^{(1)})}{p_0(\bar{z}_0)} \exp \left\{ f(\bar{z}_0 - z_0^{(2)} + z_0^{(1)}, 0) - f(\bar{z}_T - z_T^{(1)} + z_T^{(2)}, T) + \int_0^T \Phi(\bar{z} - z^{(2)} + z^{(1)}) dt - f(\bar{z}_0, 0) + f(\bar{z}_T, T) - \int_0^T \Phi(\bar{z}) dt - \frac{1}{2} \int_0^T \dot{z}^{(1)2} dt + \frac{1}{2} \int_0^T \dot{z}^{(2)2} dt + \int_0^T [\dot{z}^{(2)} - \dot{z}^{(1)}] d[\bar{z} - z^{(2)}] \right\},$$

where

$$\Phi(z) = - \frac{1}{2} m^2(z, t) - \frac{1}{2} \frac{\partial m(z, t)}{\partial z} + \frac{\partial f(z, t)}{\partial t}.$$

If in (29) one replaces the functions $\Phi(\bar{z} - z^{(2)} + z^{(1)})$, $\Phi(\bar{z})$, $f(\bar{z}_0 - z_0^{(2)} + z_0^{(1)}) \ln p_0(\bar{z}_0)$, ... by the functions $\Phi(z^{(1)})$, $\Phi(z^{(2)})$, $f(z_0^{(1)})$, $\ln p_0(z_0^{(2)})$, ..., expression (29) becomes $W[z^{(1)}] / W[z^{(2)}]$. Let us find an

upper estimate for the difference between the first group of functions and the second one. Since these functions are uniformly continuous in z , this difference can be made arbitrarily small (for ϵ sufficiently small) simultaneously for all trajectories $\tilde{z}(t) \in \Delta_\epsilon$. As an example let us take the integral

$$\int_0^T |\Phi(\tilde{z} - z^{(2)} + z^{(1)}) - \Phi(z^{(1)})| dt$$

and let us show that it can be made smaller than any $\delta > 0$. As a consequence of the uniform continuity we can find a μ such that

$$|\Phi(z - z^{(2)} + z^{(1)}) - \Phi(z^{(1)})| < \frac{\delta}{T} \quad \text{for} \quad |\tilde{z} - z^{(2)}| < \mu, \quad 0 \leq t \leq T.$$

Choosing $\epsilon = \mu/M_h$, we have

$$(30) \quad \int_0^T |\Phi(\tilde{z} - z^{(2)} + z^{(1)}) - \Phi(z^{(1)})| dt < \delta; \quad \tilde{z}(t) \in \Delta_\epsilon.$$

Therefore one can prove a limiting relation of the type (15) for the sum of all the terms in the exponent that give the difference between (29) and $W[z^{(1)}]/W[z^{(2)}]$, whence

$$\lim_{\epsilon \rightarrow 0} \frac{\mu_z[\Lambda_\epsilon]}{\mu_z[\Lambda_\epsilon]} = \frac{W[z^{(1)}]}{W[z^{(2)}]}.$$

For the special case when m is independent of t , when there exists a stationary distribution, and when it is taken as $p_0(z_0)$, the functional (26) is given by a formula that is symmetric in the sign of time:

$$W[z(t)] = \exp \left\{ -\frac{1}{2} \int_0^T \left[z^2 + m^2(z) + \frac{\partial m(z)}{\partial z} \right] dt - f(z_0) - f(z_T) \right\},$$

which was found in [2].

Note that by virtue of the equality $df - (\partial f/\partial t) dt = (\partial f/\partial z) dz$ formula (26) may be written in the form

$$(31) \quad W[z(t)] = p_0(z_0) \exp \left\{ -\frac{1}{2} \int_0^T \left[(\dot{z} - m)^2 + \frac{\partial m}{\partial z} \right] dt \right\}.$$

3. Consider now the process $x(t)$ corresponding to the equation

$$(32) \quad \frac{\partial p_x}{\partial t} = \frac{1}{2} \frac{\partial}{\partial x} \left[b(x, t) \frac{\partial p_x}{\partial x} \right],$$

where the function $b(x, t) = \sigma^2(x, t)$ is differentiable once with respect to t and twice with respect to x , and fulfills the conditions

It is a feature of equation (31) $\partial/\partial x = (dt/2) (\partial/\partial x) b(x, t) \partial/\partial x$ is symmetric the first and second kind coincide

Let $x(t)$ satisfy the initial condition

Define the probability function

$x(t) \in B$, satisfying the same condition

$$(33) \quad \frac{W[x^{(1)}(t)]}{W[x^{(2)}(t)]} = \lim_{\epsilon \rightarrow 0} \frac{P\{x^{(1)} \in B\}}{P\{x^{(2)} \in B\}}$$

when this limit exists and is independent of ϵ

the ratio in the right-hand side of (33)

$$(34) \quad z(t) = \int_0^t \sigma(z, t) dt$$

By virtue of the conditions imposed on σ and takes the function $x(t) \in B$ for $t \in B$. Clearly

$$(35) \quad P \left\{ x^{(1)} < x < x^{(2)} \right\} = P \left\{ z^{(1)} < z < z^{(2)} \right\}$$

where

$$z^{(i)}(t) = Z(x^{(i)}(t), t); \quad z^{(i)}(0) = x^{(i)}$$

and therefore $h_i(t) = h(t) + O(\epsilon)$.

After the usual calculations relating to the probability density $p_z = p_x dx/dz$ satisfies the equation

$$(36) \quad \frac{\partial p_z}{\partial t} = \frac{\partial}{\partial x} \left\{ \left[\frac{1}{2} \frac{\partial \ln b}{\partial x} \right] p_z \right\}$$

If we now apply the results of [2]

$$(37) \quad \lim_{\epsilon \rightarrow 0} \frac{P\{z^{(1)} < z < z^{(2)}\}}{P\{z^{(2)} < z < z^{(1)}\}} = 1$$

$$0 < \delta < b(x, t) < L.$$

It is a feature of equation (31) that its infinitesimal operator $\mathcal{L} = (dt/2) (\partial/\partial x) b(x, t) \partial/\partial x$ is symmetric, so that the Kolmogorov equations of the first and second kind coincide in form.

Let $x(t)$ satisfy the initial condition $x(0) = x_0$.

Define the probability functional $W[x(t)]$ on the space of functions $x(t) \in B$, satisfying the same condition $x(0) = x_0$, by means of the equality

$$(33) \quad \frac{W[x^{(1)}(t)]}{W[x^{(2)}(t)]} = \lim_{\epsilon \rightarrow 0} \frac{P\{x^{(1)} < x < x^{(1)} + \epsilon h \sigma(x^{(1)}), 0 \leq t \leq T\}}{P\{x^{(2)} < x < x^{(2)} + \epsilon h \sigma(x^{(2)}), 0 \leq t \leq T\}},$$

when this limit exists and is independent of $h(t) > 0$ of B . In order to calculate the ratio in the right-hand side of (33) let us make the change of variables

$$(34) \quad z(t) = \int_0^{x(t)} \frac{dx'}{\sigma(x', t)} \equiv Z(x(t), t).$$

By virtue of the conditions imposed on $b(x, t)$, this transform always exists and takes the function $x(t) \in B$ into a function $z(t)$ belonging to the same space B . Clearly

$$(35) \quad \begin{aligned} & P \left\{ x^{(i)} < x < x^{(i)} + \epsilon h \sigma(x^{(i)}), 0 \leq t \leq T \right\} \\ &= P \left\{ z^{(i)} < z < z^{(i)} + \epsilon h_i, 0 \leq t \leq T \right\}, \quad (i = \bar{1}, 2), \end{aligned}$$

where

$$z^{(i)}(t) = Z(x^{(i)}(t), t); \quad z^{(i)}(t) + h_i(t) = Z(x^{(i)}(t) + \epsilon h(t) \sigma(x^{(i)}(t)), t)$$

and therefore $h_i(t) = h(t) + O(\epsilon)$.

After the usual calculations related to the change of variable, we find that the probability density $p_z = p_x dx/dz$ corresponding to the new process $z(t)$ satisfies the equation

$$(36) \quad \frac{\partial p_z}{\partial t} = \frac{\partial}{\partial x} \left\{ \left[\frac{1}{2} \frac{\partial \ln \sigma}{\partial z} + \frac{\partial Z(x, t)}{\partial t} \right] p_z \right\} + \frac{1}{2} \frac{\partial^2 p}{\partial z^2}.$$

If we now apply the results of the previous section, we obtain

$$(37) \quad \lim_{\epsilon \rightarrow 0} \frac{P\{z^{(1)} < z < z^{(1)} + \epsilon h, 0 \leq t \leq T\}}{P\{z^{(2)} < z < z^{(2)} + \epsilon h, 0 \leq t \leq T\}} = \frac{W_z[z^{(1)}]}{W_z[z^{(2)}]},$$

$$W_z[z] = \exp \left\{ -\frac{1}{2} \int_0^T \left[\left(z - \frac{1}{2\sigma} \frac{\partial \sigma}{\partial z} - \frac{\partial Z(x, t)}{\partial t} \right)^2 + \frac{1}{2} \frac{\partial^2 \ln \sigma}{\partial z^2} + \frac{\partial}{\partial z} \left(\frac{\partial Z(x, t)}{\partial t} \right) \right] dt \right\}.$$

By virtue of (33), (35), we have

$$W[x(t)] = \exp \left\{ -\frac{1}{2} \int_0^T \left[\left(\dot{x} - \frac{1}{2} \frac{\partial \sigma}{\partial x} - \frac{\partial Z(x, t)}{\partial t} \right)^2 + \frac{1}{2} \sigma \frac{\partial^2 \sigma}{\partial x^2} + \sigma \frac{\partial^2 Z(x, t)}{\partial x \partial t} \right] dt \right\}$$

where we must set $z(t) = Z(x(t), t)$, i. e.

$$\begin{aligned} W[x(t)] &= \exp -\frac{1}{2} \int_0^T \left[\left(\dot{x} - \frac{1}{2} \frac{\partial \sigma}{\partial x} \right)^2 + \sigma \frac{\partial}{\partial t} \left(\frac{1}{\sigma} \right) dt \right] \\ &= \left(\frac{\sigma(x_T, T)}{\sigma(x_0, 0)} \right)^{\frac{1}{2}} \exp \left\{ -\frac{1}{2} \int_0^T \left[\frac{\dot{x}^2}{\sigma^2} + \frac{1}{4} \left(\frac{\partial \sigma}{\partial x} \right)^2 + \frac{1}{2} \sigma \frac{\partial^2 \sigma}{\partial x^2} \right] dt \right\} \\ (38) \quad &= \left[\frac{b(x_T, T)}{b(x_0, 0)} \right]^{\frac{1}{4}} \exp \left\{ -\frac{1}{2} \int_0^T \left[\frac{\dot{x}^2}{b} - \frac{1}{16} \frac{1}{b} \left(\frac{\partial b}{\partial x} \right)^2 + \frac{1}{4} \frac{\partial^2 b}{\partial x^2} \right] dt \right\} \end{aligned}$$

4. Finally let us consider the general case of a one-dimensional diffusion process $x(t)$ having the infinitesimal operator

$$(39) \quad dL(t) = \frac{1}{2} b(x, t) \frac{\partial^2}{\partial x^2} + a(x, t) \frac{\partial}{\partial x},$$

where $b(x, t), a(x, t)$ are functions with properties similar to those mentioned above.

If we set its transition probability equal to

$$p_{t_1, t}(x_1, x) = e^{f(x_1, t)} \tilde{p}_{t_1, t}(x_1, x) e^{-f(x, t)},$$

the infinitesimal operator

$$(40) \quad d\tilde{L} = e^{-f(x, t)} dL e^{f(x, t)} + \frac{\partial f(x, t)}{\partial t} dt,$$

will correspond to the function $p_{t_1, t}(x_1, x)$, or, substituting (39)

$$(41) \quad \frac{d\tilde{L}}{dt} = \frac{1}{2} \frac{\partial}{\partial x} b \frac{\partial}{\partial x} + \left[b \frac{\partial f}{\partial x} - \frac{1}{2} \frac{\partial b}{\partial x} + a \right] \frac{\partial}{\partial x} + a \frac{\partial f}{\partial x} + \frac{1}{2} b \left[\left(\frac{\partial f}{\partial x} \right)^2 + \frac{\partial^2 f}{\partial x^2} \right] + \frac{\partial f}{\partial t}$$

To make the term with the first

$$\frac{\partial f}{\partial t} =$$

By analogy with §2, one can prove with respect to the auxiliary process $\tilde{x}(t)$ that $(\partial/\partial x) b \partial/\partial x$. By virtue of (39)

$$\begin{aligned} \frac{d\mu_z[x|x_0]}{d\mu_z[x|x_0]} &= \exp \left\{ f(x_0, 0) \right. \\ &\quad \left. + \int_0^T a \frac{\partial f}{\partial x} dt + \right\} \end{aligned}$$

The probability functional $W[x(t)]$ is the distribution density $(p_0(x_0))$ multiplied by the indicated auxiliary process $x(t)$

$$W[x(t)] = \sqrt{\sigma(x_0, 0) \sigma(x_T, T)} p_0(x_0)$$

$$\int_0^T \left[\frac{\dot{x}^2}{\sigma^2} + \frac{1}{4} \left(\frac{\partial \sigma}{\partial x} \right)^2 + \frac{1}{2} \sigma \frac{\partial^2 \sigma}{\partial x^2} \right] dt$$

Substituting (42), we transform this

$$W[x(t)] = \sqrt{\sigma(x_0, 0) \sigma(x_T, T)} p_0(x_0)$$

$$-\frac{1}{2} \int_0^T \left[\frac{a^2}{b} + b \frac{\partial}{\partial x} \left(\frac{a}{b} \right) - \frac{1}{4} b \right] dt$$

Introducing the notation $m = \sigma^{-1}(x)$ the functional in the following form:

$$W[x(t)] = \sigma(x_0, 0) p_0(x_0) \exp \left\{ -\frac{1}{2} \int_0^T \left[\frac{a^2}{b} + b \frac{\partial}{\partial x} \left(\frac{a}{b} \right) - \frac{1}{4} b \right] dt \right\}$$

This expression coincides with (31) if we substitute $p_0(x_0)$ by $\tilde{z}, \partial m/\partial z, p_0(z_0)$. It

To make the term with the first derivative vanish, let us set

$$\frac{\partial f}{\partial t} = \frac{1}{2b} \frac{\partial b}{\partial x} - \frac{a}{b}.$$

By analogy with §2, one can prove that the process $x(t)$ is absolutely continuous with respect to the auxiliary process $\tilde{x}(t)$ determined by the infinitesimal operator $\frac{1}{2}(\partial/\partial x) b \partial/\partial x$. By virtue of (39)–(42) the functional derivative is equal to

$$\frac{d\mu_z[x|x_0]}{d\mu_{\tilde{x}}[x|x_0]} = \exp \left\{ f(x_0, 0) - f(x_T, T) + \int_0^T \frac{\partial f}{\partial t} dt + \int_0^T a \frac{\partial f}{\partial x} dt + \frac{1}{2} \int_0^T b \left[\left(\frac{\partial f}{\partial x} \right)^2 + \frac{\partial^2 f}{\partial x^2} \right] dt \right\}.$$

The probability functional $W[x(t)]$ is defined as the product of the initial distribution density $(p_0(x_0))$ multiplied by $\sigma(x_0, 0)$, the probability functional of the indicated auxiliary process $x(t)$ (38), and the functional (43):

$$W[x(t)] = \sqrt{\sigma(x_0, 0) \sigma(x_T, T)} p_0(x_0) \exp \left\{ f(x_0, 0) - f(x_T, T) + \int_0^T \frac{\partial f}{\partial t} dt - \frac{1}{2} \int_0^T \left[\frac{\dot{x}^2}{\sigma^2} + \frac{1}{4} \left(\frac{\partial \sigma}{\partial x} \right)^2 + \frac{1}{2} \sigma \frac{\partial^2 \sigma}{\partial x^2} \right] dt + \int_0^T a \frac{\partial f}{\partial x} dt + \frac{1}{2} \int_0^T b \left[\left(\frac{\partial f}{\partial x} \right)^2 + \frac{\partial^2 f}{\partial x^2} \right] dt \right\}.$$

Substituting (42), we transform this expression into the form

$$W[x(t)] = \sqrt{\sigma(x_0, 0) \sigma(x_T, T)} p_0(x_0) \exp \left\{ -\frac{1}{2} \int_0^T \frac{\dot{x}^2}{b} dt + \int_0^T \left[a - \frac{1}{2} \frac{\partial b}{\partial x} \right] \frac{\dot{x}}{b} dt - \frac{1}{2} \int_0^T \left[\frac{a^2}{b} + b \frac{\partial}{\partial x} \left(\frac{a}{b} \right) - \frac{1}{4} b \frac{\partial}{\partial x} \left(\frac{1}{b} \frac{\partial b}{\partial x} \right) - \frac{1}{16} \frac{1}{b} \left(\frac{\partial b}{\partial x} \right)^2 \right] dt \right\}.$$

Introducing the notation $m = \sigma^{-1}(a - \frac{1}{4}(\partial b/\partial x))$, we can write the probability functional in the following form:

$$W[x(t)] = \sigma(x_0, 0) p_0(x_0) \exp \left\{ -\frac{1}{2} \int_0^T \left[\frac{\dot{x}^2}{b} - 2 \frac{m\dot{x}}{\sigma} + m^2 + \sigma \frac{\partial m}{\partial x} \right] dt \right\}.$$

This expression coincides with (31) if we replace $\dot{x}/\sigma, \sigma \partial m/\partial x, p_0(x_0)$ by $\dot{z}, \partial m/\partial z, p_0(z_0)$. It follows from this that Theorem 1

